

# Recovery Theory

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Fallstudien der mathematischen Modellbildung, Teil 2

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- $y = Ax_0 + e$ ,  $\|e\| \leq \delta$ ,  $A \in \mathbb{R}^{m \times n}$   $m < n$ . Assume  $x_0$  is  $k$ -sparse, and consider the LASSO

$$x = \operatorname{argmin}_x \frac{1}{2} \|y - Ax\|^2 + \lambda \|x\|_1$$

- Questions: is  $x = x_0$  when  $\delta = 0$ ,  $\lambda = 0$ ? how close is  $x$  from  $x_0$ ?  
→ depends on properties of  $A$ , and on  $k$ .

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- Noiseless, constrained problem **basis pursuit**

$$\min \|x\|_1 \quad \text{s.t.} \quad Ax = y$$

We will always assume that the problem is feasible, *i.e.* there exists  $x_0 \in \mathbb{R}^n$  such that  $Ax_0 = y$

- For  $f_0$  is differentiable, consider

$$\min f_0(x) \quad \text{s.t.} \quad Ax = y$$

Is  $x_0$  a minimizer? The (sufficient) KKT conditions for optimality give

$$\begin{cases} \nabla f_0(x) - A^T \nu = 0 \\ Ax - y = 0 \end{cases}$$

Therefore, if there exists  $\nu$  such that  $\nabla f_0(x_0) = A^T \nu$ , then  $x_0$  is a solution.  $A^T \nu$  is called a *dual certificate*. In that case,  $\nu$  is also a solution of the dual problem. The dual certificate is not necessary unique.

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- For basis pursuit, KKT at  $x_0$  gives

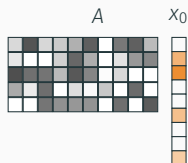
$$\begin{cases} A^T \nu \in \partial \|x\|_1 \\ Ax - y = 0 \end{cases}$$

They are also sufficient: if  $A^T \nu \in \partial \|x_0\|_1$ , then for all  $x$

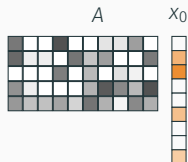
$$\|x\|_1 \geq \|x_0\|_1 + \langle A^T \nu, x - x_0 \rangle \iff \|x\|_1 \geq \|x_0\|_1 + \langle \nu, Ax - Ax_0 \rangle$$

which directly shows that  $\|x\|_1 \geq \|x_0\|_1$  if  $x$  is feasible.

- Let  $I = \text{Supp } x_0$ ,  $I^c$  its complement,  $x_I$  the restriction of  $x \in \mathbb{R}^n$  to  $I$  and  $A_I$  the column restriction of  $A$  to  $I$

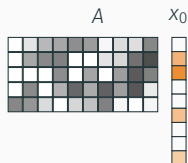


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- $\eta = A^T \nu \in \partial \|x_0\|_1$  when
  - $\eta_I = \text{sign}(x_0)_I$
  - $\|\eta_{I^c}\|_\infty \leq 1$

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- Theorem (Fuchs, 2004).**  $x_0$  is the unique minimizer of

$$\min \|x\|_1 \quad \text{s.t.} \quad Ax = Ax_0$$

if there exists  $\eta \in \text{Ran } A^T$  such that

- $\eta_I = \text{sign}(x_0)_I$
- $\|\eta_{I^c}\|_\infty < 1$  ("non-degeneracy")

and furthermore  $A$  is injective on  $I$ .



Let  $x \in \mathbb{R}^n$  such that  $Ax = Ax_0$ . We have

$$\begin{aligned}\|x_0\|_1 &= \langle \eta, x_0 \rangle \\ &= \langle \eta, x \rangle \\ &= \langle \eta_I, x_I \rangle + \langle \eta_{I^c}, x_{I^c} \rangle \\ &\leq \|x_I\|_1 + \|\eta_{I^c}\|_\infty \|x_{I^c}\|_1\end{aligned}$$

If  $x_{I^c} \neq 0$ , then  $\|x_0\|_1 < \|x\|_1$

If  $x_{I^c} = 0$  (i.e.  $x$  has the same support as  $x_0$ ), then  $A(x - x_0) = A_I(x_I - (x_0)_I) = 0$ , so by injectivity  $x_I = (x_0)_I$ , and therefore  $x = x_0$ .

Hence,  $\|x_0\|_1 < \|x\|_1$  unless  $x = x_0$ .

□

- Classical way: look for **minimal norm certificate**

$$\min \|\nu\|_2^2 \quad \text{s.t.} \quad (A^\top \nu)_I = \text{sign}(x_0)_I$$

- Note that  $(A^\top \nu)_I = \text{sign}(x_0)_I \iff (A_I)^\top \nu = \text{sign}(x_0)_I$ : underdetermined system for sufficiently sparse vector

The minimal norm solution is given by the pseudo-inverse, *i.e.*

$$\eta_0 = A^\top (A_I^\top)^\dagger \text{sign}(x_0)_I$$

Since  $A_I$  is injective,  $\eta_I = A_I^\top (A_I)^\dagger \text{sign}(x_0)_I = \text{sign}(x_0)_I$

- It remains to show that  $\|\eta_{I^c}\|_\infty < 1$

## EXAMPLE: GAUSSIAN RANDOM MATRIX

- Assume  $A_{ij} \sim \mathcal{N}(0, 1)$  i.i.d.. This means in particular

$$\mathbb{P}(|A_{ij}| \geq u) \propto 2 \int_u^\infty \exp\left(-\frac{t^2}{2}\right) dt \leq \exp\left(-\frac{u^2}{2}\right)$$
$$\int_u^\infty e^{-t^2/2} dt = \int_0^\infty e^{-(t+u)^2/2} dt = e^{-u^2/2} \int_0^\infty e^{-tu} e^{-t^2/2} dt$$
$$\begin{cases} \leq e^{-u^2/2} \int_0^\infty e^{-t^2/2} dt = \sqrt{\frac{\pi}{2}} e^{-u^2/2} \\ \leq e^{-u^2/2} \int_0^\infty e^{-tu} dt \leq \frac{1}{u} e^{-u^2/2} \end{cases}$$

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$$\eta_i = \sum_{j=1}^m A_{ji} \left[ (A_l^\top)^\dagger \text{sign}(x_0)_l \right]_j =: \sum A_{ji} c_j$$

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- Lemma (1).** If  $X_i \sim \mathcal{N}(0, 1)$  i.i.d., then  $\mathbb{P}(|\sum c_i X_i| \geq u) \leq \exp\left(-\frac{u^2}{2\|c\|_2^2}\right)$

**Lemma (2, admitted).** If  $A_{ij} \sim \mathcal{N}(0, 1)$  i.i.d.,  $i = 1, \dots, m, j = 1, \dots, k, k < m$ , then

$$\mathbb{P}\left(\sigma_{\max}\left(\frac{A}{\sqrt{m}}\right) \geq 1 + \sqrt{\frac{k}{m}} + t\right) \leq \exp\left(-\frac{mt^2}{2}\right)$$

$$\mathbb{P}\left(\sigma_{\min}\left(\frac{A}{\sqrt{m}}\right) \leq 1 - \sqrt{\frac{k}{m}} + t\right) \leq \exp\left(-\frac{mt^2}{2}\right)$$

Estimate  $P = \mathbb{P}(\exists i \notin I : |\eta_i| \geq 1)$ . Consider the event  $E: \|c\|_2 \leq \alpha$  for some  $\alpha > 0$ .

■ Then

$$\begin{aligned} P &= \mathbb{P}(\exists i \notin I : |\eta_i| \geq 1 | E) \mathbb{P}(E) + \mathbb{P}(\exists i \notin I : |\eta_i| \geq 1 | E^c) \mathbb{P}(E^c) \\ &\leq \mathbb{P}(\exists i \notin I : |\eta_i| \geq 1 | E) + \mathbb{P}(E^c) \end{aligned}$$

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- For a fixed  $i \notin I$ , using Lemma 1,

$$\mathbb{P}(|\sum A_{ij}c_j| \geq 1 | E) \leq \exp\left(-\frac{1}{2\|c\|_2^2}\right) \leq \exp\left(-\frac{1}{2\alpha^2}\right),$$

and since there are  $(n - k)$  possible values of  $i \notin I$ , we obtain

$$\mathbb{P}(\exists i \notin I : |\eta_i| \geq 1 | E) \leq (n - k) \exp\left(-\frac{1}{2\alpha^2}\right)$$

## FAILURE PROBABILITY (1/2)

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$$\|c\|_2 = \|(A_I)^\dagger \text{sign}(x_0)_I\|_2 \leq (\sigma_{\min}(A_I))^{-1} \sqrt{k} = \sigma_{\min}^{-1}(A_I / \sqrt{k})$$



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and hence

$$\begin{aligned} \mathbb{P}(E^c) &\leq \mathbb{P}(\sigma_{\min}^{-1}(A_I / \sqrt{k}) \geq \alpha) = \mathbb{P}\left(\sigma_{\min}(A_I / \sqrt{m}) \leq \frac{1}{\alpha} \sqrt{\frac{k}{m}}\right) \\ &\leq \exp\left(-\frac{m}{2} \left(1 - \left(1 + \frac{1}{\alpha}\right) \sqrt{\frac{k}{m}}\right)^2\right) \end{aligned}$$

Let  $\varepsilon > 0$ , we want  $(a) + (b) \leq \varepsilon$

- Set

$$\exp\left(-\frac{1}{2\alpha^2}\right) = \frac{\varepsilon}{n}$$

*i.e.*  $\alpha^{-2} = 2 \ln(n/\varepsilon)$ , then  $(a) = (n - k)\varepsilon/n \leq \frac{n-1}{n}\varepsilon$

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- We would like  $(b) \leq \varepsilon/n$

$$\iff m \left(1 + \frac{1}{\alpha}\right) \sqrt{\frac{k}{m}} \geq 2 \ln\left(\frac{n}{\varepsilon}\right) = \alpha^{-2}$$

$$\iff \sqrt{m} - \left(1 + \frac{1}{\alpha}\right) \sqrt{k} \geq \alpha^{-1}$$

$$\iff \sqrt{m} \geq \alpha^{-1} + (1 + \alpha^{-1})\sqrt{k}$$

$$\iff \sqrt{m} \geq 3\alpha^{-1}\sqrt{k}$$

$$\iff m \geq 18k \ln\left(\frac{n}{\varepsilon}\right)$$

$$\iff \# \text{ measurements} \gtrsim \text{sparsity} \times \ln(n)$$

**Theorem (Donoho; Candès-Romberg-Tao).** Let  $x_0 \in \mathbb{R}^n$  be  $k$ -sparse,  $y = Ax_0$  for  $A \in \mathbb{R}^{m \times n}$ ,  $A_{ij} \sim \mathcal{N}(0, 1)$  i.i.d. Then

$$\min \|x\|_1 \quad \text{s.t.} \quad Ax = y$$

recovers  $x_0$  with probability  $\geq 1 - \epsilon$ , provided

$$m \geq Ck \ln \left( \frac{n}{\epsilon} \right)$$

*Remark (Tighter analysis).* A sharper condition can be derived

$$m \geq 2k \ln \left( \frac{en}{k} \right) - \ln \epsilon$$

in general for sub-gaussian matrices.

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- **Definition (Best  $k$ -term approximation).** For  $p > 0$ ,  $x \in \mathbb{R}^n$ ,

$$\sigma_k(x)_p := \inf \{ \|x - y\|_p ; y \text{ is } k\text{-sparse} \}$$

$\inf$  is realised by taking  $y = x_I$  where  $I$  is the support of the  $k$  largest entries of  $x$

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  - $\|w\|_2 \leq \varepsilon$
  - injectivity of  $A$  controlled on  $I$  with  $|I| = k$then any recovered  $x$  obeys

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- Let  $A = [a_1, \dots, a_n]$  and

$$x = \operatorname{argmin} \|x\|_1 \quad \text{s.t.} \quad \|Ax - y\| \leq \varepsilon$$

**Proposition (Dual Certification for Inexact Data<sup>1</sup>).** Let  $x_0 \in \mathbb{R}^n$  with  $k$  largest components on  $I$ . Let  $y = Ax_0 + w$  with  $\|w\| \leq \varepsilon$ . Assume  $\exists \delta, \beta, \gamma, \theta, \tau$  ( $\delta < 1$ ,  $\theta < 1$ ) and  $\eta = A^T \nu$  such that

- $\|A_I^T A_I - I_n\|_{2,2} \leq \delta$
- $\|\eta_I - \operatorname{sign}(x_0)_I\|_2 \leq \gamma$
- $\|\nu\|_2 \leq \tau\sqrt{k}$ .
- $\max_{i \notin I} \|A_I^T a_i\|_2 \leq \beta$
- $\|\eta_{I^c}\|_\infty \leq \theta$

If  $\theta + \beta\gamma/(1 - \delta) < 1$ , then

$$\|x - x_0\|_2 \leq C_1 \sigma_k(x)_1 + C_2 \sqrt{k} \varepsilon$$

<sup>1</sup>A Mathematical Introduction to Compressive Sensing, 1993



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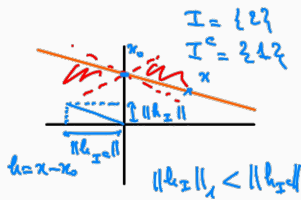
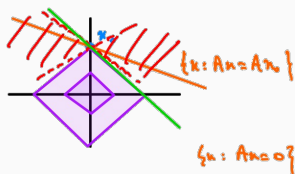
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- Recoverability of  $x_0 \iff$  favorable orientation of  $\text{Ker } A$

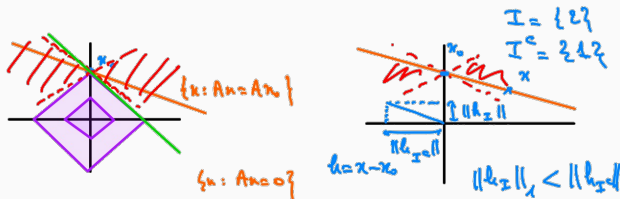


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- Definition (NSP).**  $A$  obeys the nullspace property relative to  $I$  if for all  $h \in \text{Ker } A \setminus 0$ ,

$$\|h_I\|_1 < \|h_{I^c}\|_1$$

it obeys the NSP of order  $k$  if it satisfies the NSP for every  $I$  such that  $|I| = k$ .

- Theorem (Uniform Recovery).** Every  $k$ -sparse  $x_0$  such that  $Ax_0 = y$  is the unique solution of (BASIS PURSUIT) if and only if  $A$  obeys NSP( $k$ ).

- Definition (Robust NSP).** A obeys the robust NSP of order  $k$  if  $\exists 0 < \rho < 1, \tau > 0$ , such that, for all  $I : |I| = k$ , for all  $h \in \text{Ker } A \setminus 0$

$$\|h_I\|_1 \leq \rho \|h_{I^c}\|_1 + \tau \|Ah\|_2$$

- Theorem (Uniform Robust Recovery).** If A obeys the robust NSP( $k$ ), then for every  $k$ -sparse  $x_0$  such that  $Ax_0 = y$ , any solution of (BASIS PURSUIT- $\epsilon$ ) satisfies

$$\|x - x_0\|_1 \leq 2 \frac{1 + \rho}{1 - \rho} \sigma_k(x_0)_1 + 4 \frac{\tau}{1 - \rho} \epsilon$$

- Limitation of NSP: number of measurements ( $m$ ) required has bad dependency on  $k$  (e.g.  $m \sim k^2$  for  $A = \begin{bmatrix} I & F \end{bmatrix}$ )

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- **Compressed Sensing**: stronger conditions on  $A$  to ensure uniform recovery with fewer measurements ( $m \sim k \log n$ )

**Definition (Restricted Isometry Property (RIP)).**  $A$  is said to satisfy the RIP if there exists  $\delta_k$ , such that for all  $k$ -sparse  $x$ ,

$$(1 - \delta_k) \|x\|_2^2 \leq \|Ax\|_2^2 \leq (1 + \delta_k) \|x\|_2^2$$

$$\delta_k = \max_{|S|=k} \|A_S^\top A_S - I_k\|_{2,2}$$